

## Lecture 11

1.

### The Basic Numerical Quantities

### Associated to a Continuous $X$

In this lecture we will introduce four basic numerical quantities associated to a continuous random variable  $X$ . You will be asked to calculate those (and the cdf of  $X$ ) given  $f(x)$  on the midterms and the final.

These quantities are

1. The  $p$ -th percentile  $\eta(p)$
2. The  $\alpha$ -th critical value  $x_\alpha$ .
3. The expected value  $E(X)$  or  $\mu$ .
4. The variance  $V(X)$  or  $\sigma^2$ .

I will compute all these for  $U(0,1)$ .  
the linear distribution and  $U(a,b)$ .

# Percentiles and Critical

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## Values of Continuous Random Variables

### Percentiles

Greek letter eta

Let  $p$  be a number between 0 and 1. The  $100p$ -th percentile denoted  $\eta(p)$ , of a continuous random variable  $X$  is the unique number satisfying

$$P(X \leq \eta(p)) = p \quad (\#)$$

or  $F(\eta(p)) = p \quad (\#\#)$

So if you know  $F$  you can find  $\eta(p)$ . Roughly

$$\eta(p) = F^{-1}(p)$$

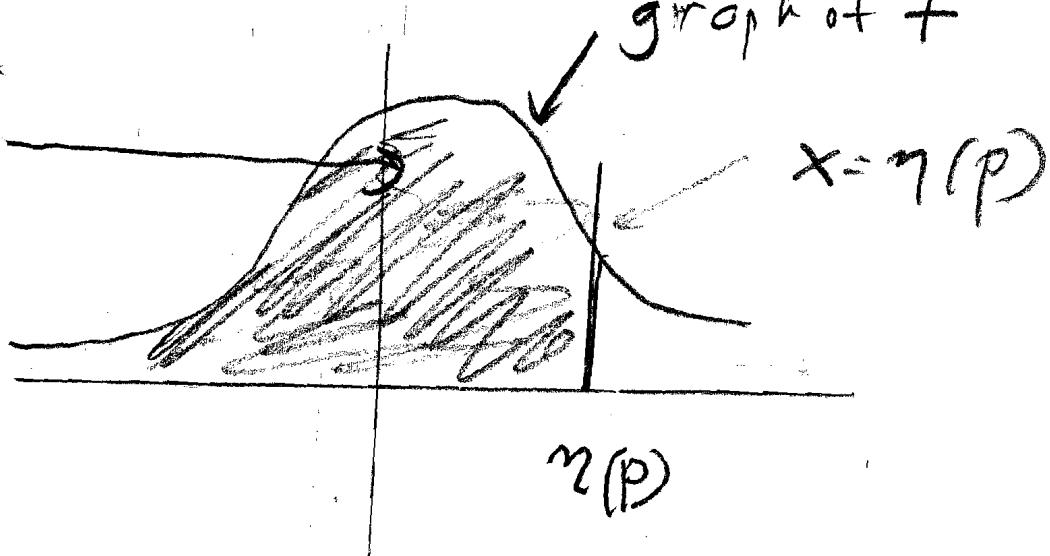
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The geometric interpretation  
of  $\gamma(p)$  is very important

graph of  $f$

This

area is  $p$



The geometric interpretation of  $(\gamma)$

$\gamma(p)$  is the number such that  
the vertical line  $x=\gamma(p)$  cuts off  
area  $p$  to the left under  
the graph of  $f(x)$ .

(this is the picture above)

## Special Case      The median $\tilde{\mu}$

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The median  $\tilde{\mu}$  is the unique number so that

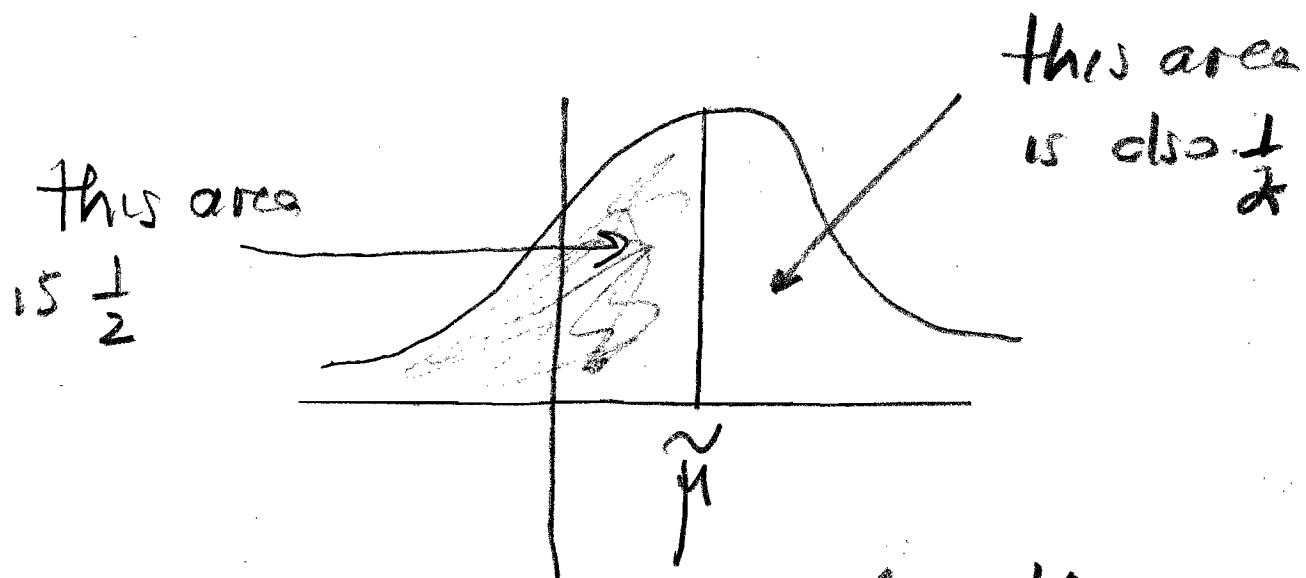
$$P(X \leq \tilde{\mu}) = \frac{1}{2}$$

or

$$F(\tilde{\mu}) = \frac{1}{2}$$

so the median is the 50-th percentile

### The picture



Since the total area is 1, the area to the right of the vertical line  $x = \tilde{\mu}$  is also  $\frac{1}{2}$ . So  $x = \tilde{\mu}$  bisects the area.

# Critical Values

Roughly speaking if you switch left to right in the definition of percentile you get the definition of the critical value | Critical values play a key role in the formulas for confidence intervals (lectr).

Definition Let  $\alpha$  be a

real number between 0 and 1.

Then the  $\alpha$ -th critical value, denoted  $x_\alpha$ , is the unique number satisfying

$$P(X \geq x_\alpha) = \alpha \quad (\text{b})$$

Let's rewrite (b) in terms  
of  $F$ . We have

$$\begin{aligned} P(X \geq x_\alpha) &= 1 - P(X \leq x_\alpha) \\ &= 1 - F(x_\alpha) \end{aligned}$$

So (b) becomes

$$1 - F(x_\alpha) = \alpha$$

$$F(x_\alpha) = 1 - \alpha$$

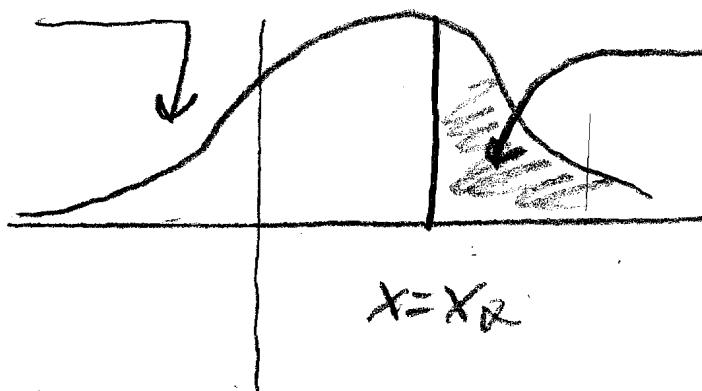
$$x_\alpha = F^{-1}(1 - \alpha) \quad (\text{bb})$$

What about the geometric  
interpretation?

# The geometric interpretation

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graph of  $f$



this area is  $\alpha$

$$x = x_\alpha$$

$x_\alpha$  is the number so that the vertical line  $x = x_\alpha$  cuts off area  $\alpha$  to the right under the graph of  $f(x)$ .

## Relation between critical values and percentiles

$x = x_\alpha$  cuts off area  $1 - \alpha$  to the left since the total area is 1. But  $\eta(1-\alpha)$  is the number such that  $x = \eta(1-\alpha)$  cuts off area  $1 - \alpha$  to the left.

So

$$x_\alpha = \eta(1-\alpha)$$

# Computation of Examples

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## Example 1

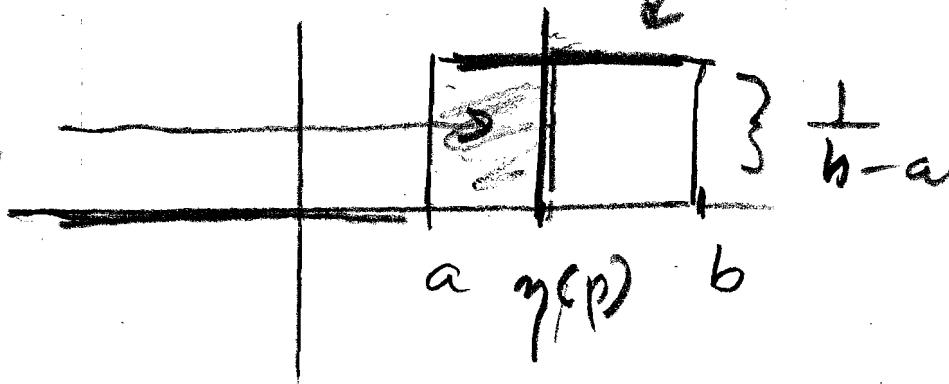
$$X \sim U(a, b)$$

Let's compute the  $\eta(p)$ -th percentile for  $X \sim U(a, b)$

graph of  $f$

This area

is  $p$



So the point  $\eta(p)$  between  $a$  and  $b$  must have the property that the area of the shaded box is  $p$ . But the base of the box is  $\eta(p) - a$  and the height is  $\frac{1}{b-a}$  so

$$\text{Area} = b \cdot h = (\eta(p) - a) \left( \frac{1}{b-a} \right) \text{ so}$$

$$(\eta(p) - a) \left( \frac{1}{b-a} \right) = p \quad \text{or}$$

$$\eta(p) = a + p(b-a) = (1-p)a + pb \quad (*)$$

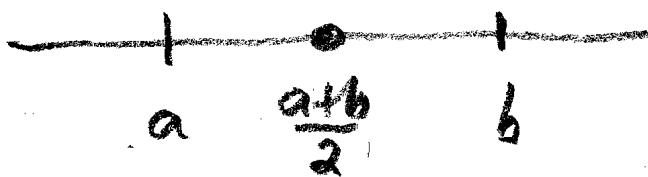
How about the median  $\tilde{\mu}$ . 9

So we want  $\gamma(\frac{1}{2})$ . By (\*)  
we have

$$\tilde{\mu} = \gamma\left(\frac{1}{2}\right) = a + \frac{b-a}{2} = \frac{a+b}{2}$$

### Remark

$\frac{a+b}{2}$  is the midpoint of  
the interval  $[a, b]$



# Critical Values for $U(a, b)$

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$$x_\alpha = \gamma(1-\alpha) = a + (1-\alpha)(b-a)$$

$$= a + b - a - \alpha b + \alpha a$$

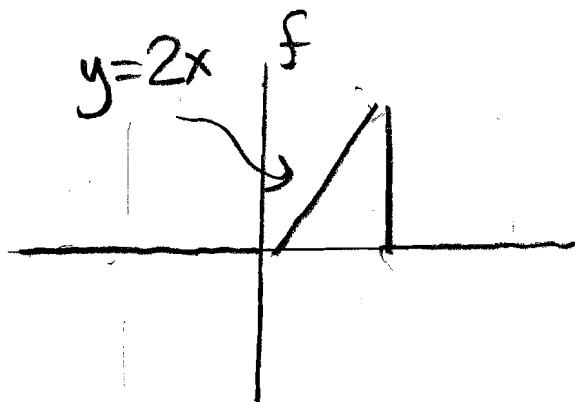
So

$$x_\alpha = \alpha a + (1-\alpha)b.$$

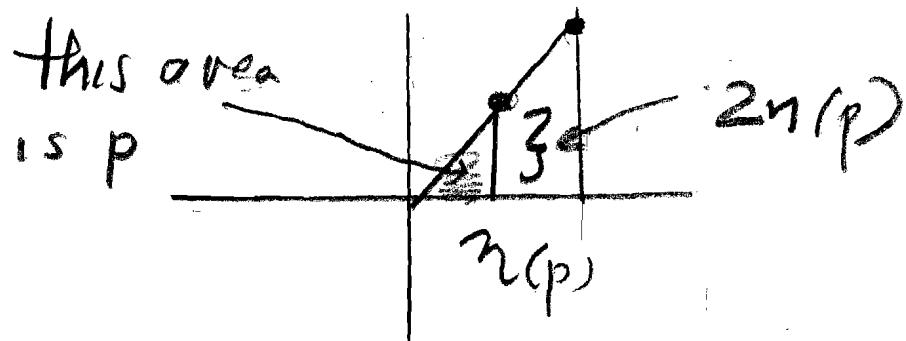
## Example 2 The linear distribution

Recall the linear distribution has density

$$f(x) = \begin{cases} 0, & x < 0 \\ 2x, & 0 \leq x \leq 1 \\ 0, & x > 1 \end{cases}$$



# The 100p-th percentile



We want the area of the triangle to be  $p$ . But the base is  $n(p)$  and the height is  $2n(p)$  so

$$\begin{aligned} A &= \frac{1}{2} b h = \frac{1}{2} n(p) (2n(p)) \\ &= n(p)^2 \end{aligned}$$

We have to solve

$$n(p)^2 = p$$

$$\text{so } n(p) = \sqrt{p} \quad \text{so}$$

In particular

$$\tilde{\mu} = \gamma\left(\frac{1}{2}\right) = \sqrt{\frac{1}{2}} = \frac{\sqrt{2}}{2}$$

This will be important later

# Expected Value

## Definition

The expected value or mean  $E(X)$  or  $\mu$  of a continuous random variable is defined by

$$E(X) = \int_{-\infty}^{\infty} x f(x) dx$$

We will compute some examples.

### Example 1

$$X \sim U(a, b)$$

$$\begin{aligned} E(X) &= \int_{-\infty}^{\infty} f(x) dx = \int_a^b \frac{1}{b-a} x dx \\ &= \frac{1}{b-a} \left( \frac{x^2}{2} \right) \Big|_{x=a}^{x=b} = \frac{1}{2} \frac{(b^2 - a^2)}{b-a} = \frac{b+a}{2} \end{aligned}$$

Now we showed on page 9  
 that if  $X \sim U(a, b)$  then the  
 median  $\tilde{\mu}$  was given by

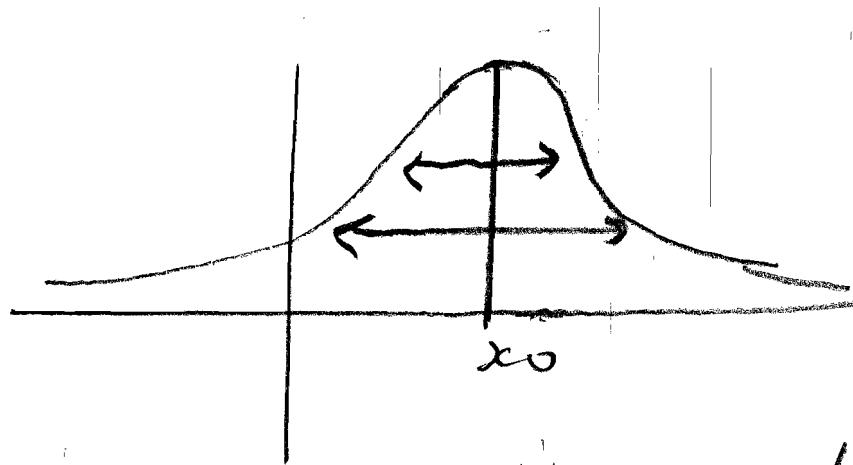
$$\tilde{\mu} = \frac{a+b}{2}$$

Hence in this the mean  
is equal to the median

$$\mu = \tilde{\mu} = \frac{a+b}{2}$$

2 This is not always the  
 case as we will see  
 shortly.

The "reason"  $\mu = \tilde{\mu}$  is that  $f(x)$  has a point of symmetry ie a point  $x_0$  so that  $f(x_0+y) = f(x_0-y)$



This means that the graph is symmetrical about the vertical line (mirror)  $x=x_0$

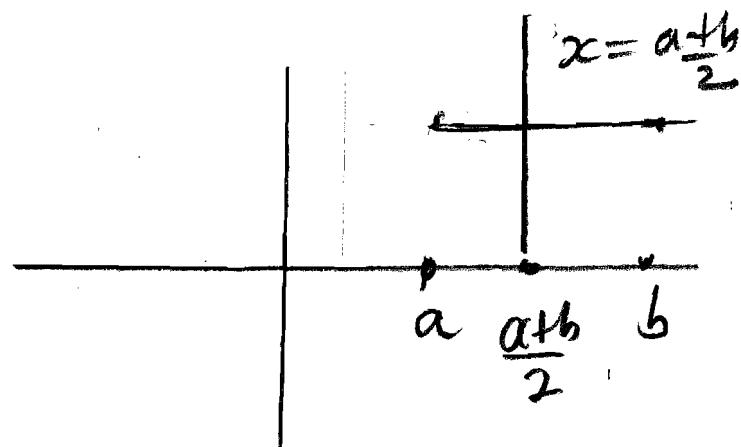
Proposition (Useful fact)

If  $x_0$  is a point of symmetry for  $f(x)$  then

$$\mu = \tilde{\mu} = x_0$$

Now if  $X \sim U(a, b)$  then

$x_0 = \frac{a+b}{2}$  is a point of symmetry  
for  $f(x)$

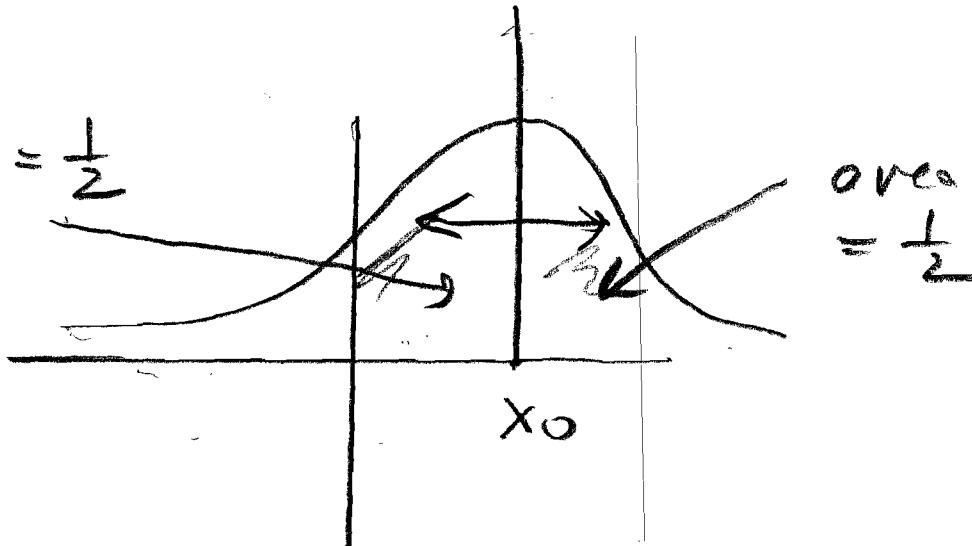


For a change we will prove  
the proposition.

Proof

$\hat{\mu} = x_0$  is immediate because  
by symmetry there is equal  
area to the left and right  
of  $x_0$ .

$$\text{area} = \frac{1}{2}$$



Since the total area is 1 the area to the left of  $x_0$  is  $\frac{1}{2}$ .  
Hence  $\hat{\mu} = x_0$ .

It is harder to prove

$$E(X) = \int_{-\infty}^{\infty} x f(x) dx = x_0$$

Trick: Since  $x_0$  is a constant and  $\int_{-\infty}^{\infty} f(x) dx = 1$  we have

$$\int_{-\infty}^{\infty} x_0 f(x) dx = x_0$$

Thus to show

$$\int_{-\infty}^{\infty} x f(x) dx = x_0$$

It suffices to show

$$\int_{-\infty}^{\infty} x f(x) dx = \int_{-\infty}^{\infty} x_0 f(x) dx$$

or

$$\int_{-\infty}^{\infty} (x - x_0) f(x) dx = 0$$

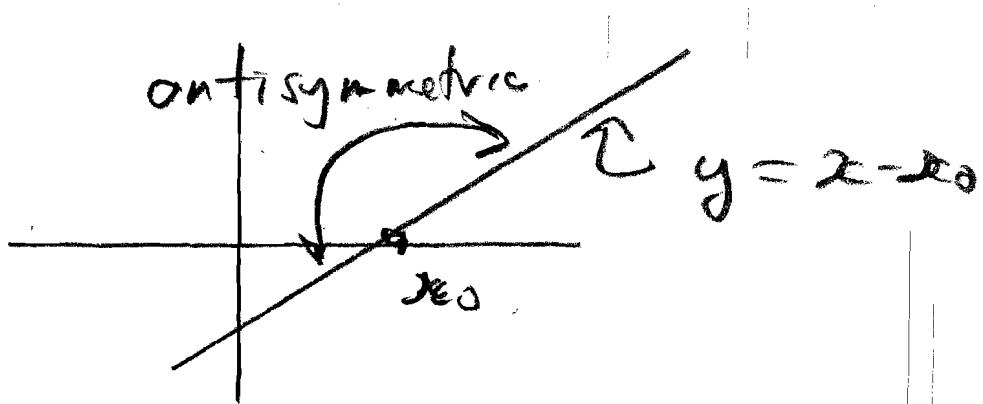
But if we put

$$g(x) = (x - x_0) f(x)$$

then  
g(x) is antisymmetric or "odd"  
about  $x_0$

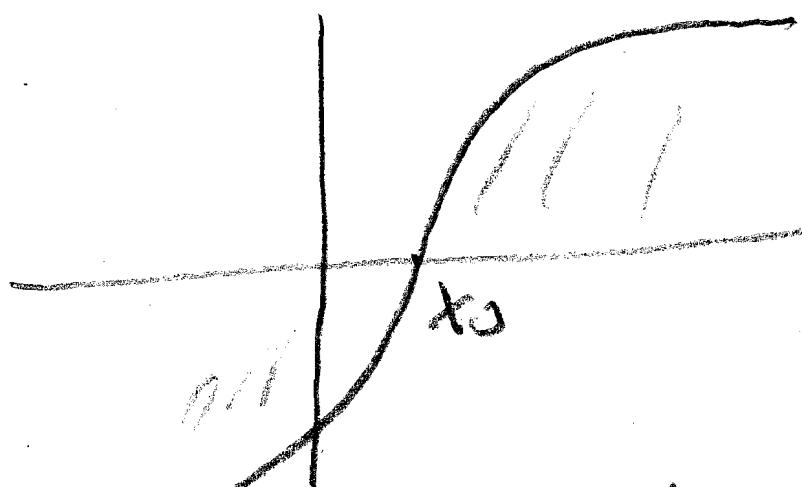
$$g(x_0 + y) = -g(x_0 - y)$$

This is because  $x - x_0$  is



But antisymmetric + symmetric  
= antisymmetric  
(or odd-even = odd)

Finally the integral of an  
antisymmetric (or "odd") function  
from  $-\infty$  to  $\infty$  is zero

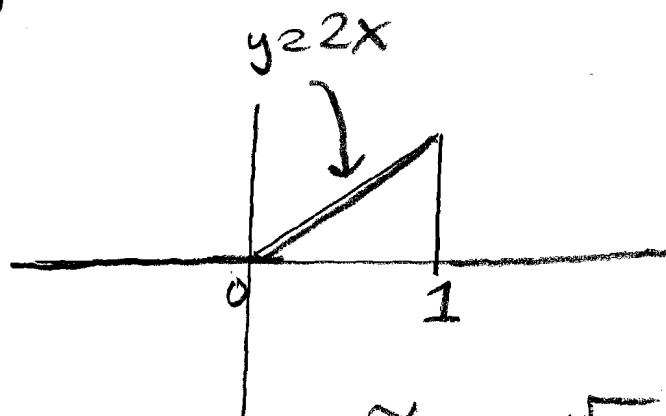


The integral to the left of  $x_0$   
cancels the area to the right

□

This fact can save a lot of parallel computation of expected values.

## Example 2 The linear distribution



We have seen  $\tilde{\mu} = \frac{\sqrt{2}}{2}$ , page 12-

$f(x)$  is certainly not symmetric so it is possible  $\mu = \tilde{\mu}$  and we will see that it is the case

$$E(X) = \int_{-\infty}^{\infty} x f(x) dx$$

$$= \int_0^1 x(2x) dx$$

$$= 2 \int_0^1 x^2 dx$$

$$= 2 \left(\frac{1}{3}\right) = \frac{2}{3}$$

Hardy fact  $\int_0^1 x^n = \frac{1}{n+1}$

$$\text{So } \mu = \frac{2}{3} \quad \text{and} \quad \tilde{\mu} = \frac{2}{5}$$

They aren't equal, which one  
is bigger?

# Variance

The Variance  $V(X)$  or  $\sigma^2$  of a continuous random variable is defined by

$$V(X) = \int_{-\infty}^{\infty} (x-\mu)^2 f(x) dx$$

Remark Once we learn about change of continuous random variable we will see

this is  $E((X-\mu)^2)$

 new random variable obtained from  $X$ .  
using  $h(x) = (x-\mu)^2$

Once again there is a  
shortcut formula for  $V(X)$ .

<u>Proposition</u>	<u>Shortcut Formula</u>
$V(X) = E(X^2) - (E(X))^2$	$= E(X^2) - \mu^2$

This is the formula to use.

Example 1  $X \sim U(a, b)$

We know  $\mu = \frac{a+b}{2}$ . We  
have to compute  $E(X^2)$

$$E(X^2) = \int_{-\infty}^{\infty} x^2 f(x) dx$$

$$= \int_a^b x^2 \frac{1}{b-a} dx$$

$$= \frac{1}{b-a} \left( \frac{x^3}{3} \right) \Big|_{x=a}^{x=b}$$

$$= \frac{1}{3} \frac{b^3 - a^3}{b-a} = \frac{1}{3} (b^2 + ab + a^2)$$

So

$$V(X) = \frac{1}{3} (a^2 + ab + b^2) - \left( \frac{a+b}{2} \right)^2$$

$$= \frac{a^2 + ab + b^2}{3} - \frac{a^2 + 2ab + b^2}{4}$$

$$= \frac{a^2 - 2ab + b^2}{12} = \frac{(b-a)^2}{12}$$

## Example 2    The linear distribution

We have seen (pg 21)

$$\mu = \frac{2}{3}$$

We need  $E(X^2)$ .

$$\begin{aligned} E(X^2) &= \int_{-\infty}^{\infty} x^2 f(x) dx \\ &= \int_{-2}^{1} x^2 (2x) dx \\ &= 2 \int_0^1 x^3 dx = 2\left(\frac{1}{4}\right) = \frac{1}{2} \end{aligned}$$

So

$$V(X) = \frac{1}{2} - \left(\frac{2}{3}\right)^2 = \frac{1}{2} - \frac{4}{9}$$

$$= \frac{9}{18} - \frac{8}{18} = \frac{1}{18}$$