

DATA FORMAT FOR A SURVIVAL STUDY

Subjects enter at random times E_i , ‘followed’ until
 $E_i + T_i = \min(E_i + X_i, E_i + C_i)$ (not both observed)
‘death-time’ ($X_i = \textit{lifetime}$), or ‘censoring time’
(e.g., $C_i = E_{\max} - E_i + \tau$ *administrative*)

Data: $\{(E_i, T_i, \Delta_i, Z_i), i = 1, \dots, n\}$ or
 $\mathcal{D} = \{(T_i, \Delta_i), i = 1, \dots, n\}$ where

$T_i =$ *time-on-test* or *event time*

$\Delta_i = I_{[X_i \leq C_i]}$ *death indicator*

Z_i *auxiliary covariates*, e.g. group indicator ξ_i ;
may be time-dependent obs on $[0, T_i)$

Objective: to estimate the marginal survival function
 $S_X(t) = P(X_1 > t) = 1 - F_X(t)$ consistently from the
data \mathcal{D} .

Assumptions: random vectors (E_i, X_i, C_i, Z_i) in-
dependent & identically distributed (*iid*), $i = 1, \dots, n$;

also (X_i, C_i) have continuous *joint density*, i.e.

$$\lim_{\delta \searrow 0} \frac{1}{\delta^2} P(X_1 \in (x, x + \delta), C_1 \in (c, c + \delta)) = f_{X,C}(x, c)$$

Lexis Diagram for an Illustrative Clinical Trial

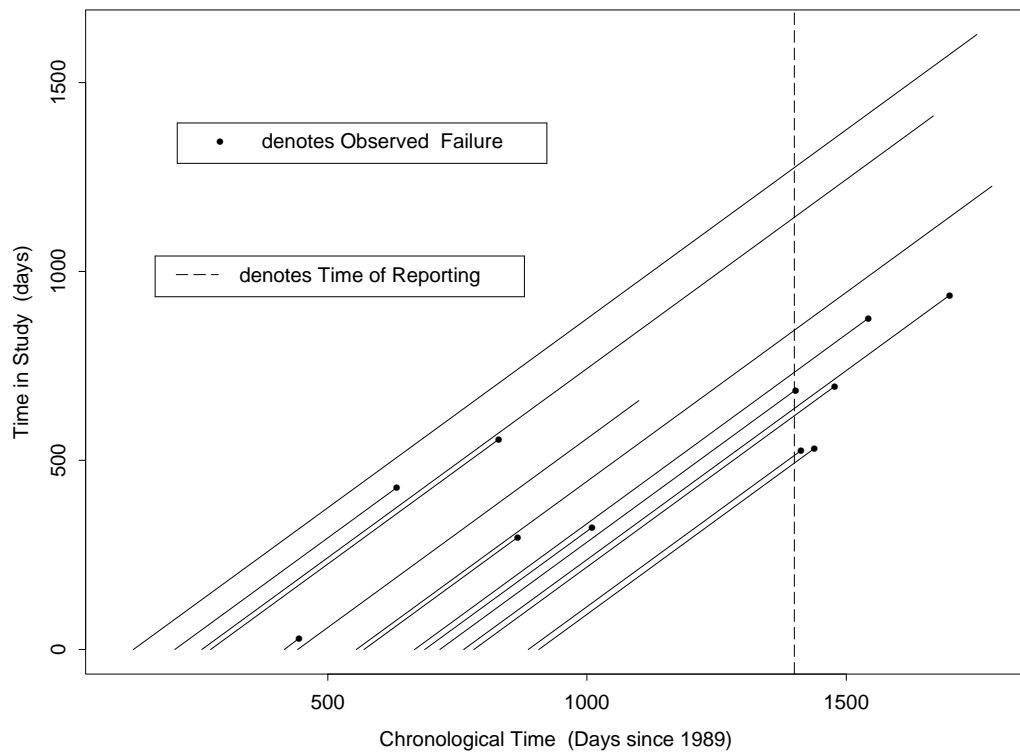


Figure 1: “Lexis Diagram” (from contributed article to Encyclopedia of Biostatistics): from entry, patients’ followup is pictured as 45° line: solid dot represents death, line not ending in dot represents censoring.

Death Hazards

In general, define **hazard intensity**

$$h_X(t) \equiv \lim_{\delta \rightarrow 0} \frac{1}{\delta} P(X \in (t, t + \delta) | X > t) = \frac{f_X(t)}{S_X(t)}$$

Then

$$h_X(t) = -\frac{d}{dt} \ln S_X(t) \Rightarrow S_X(t) = \exp\left(-\int_0^t h_X(s) ds\right)$$

So **hazard** is instantaneous mortality rate conditional on previous survival, and the integrated form of **cumulative hazard**

$$H_X(t) = \int_0^t h_X(s) ds = -\ln S_X(t)$$

is also very useful in specifying survival models.

MAJOR CASES:

(i) *Constant hazard rate:* $h_X(t) \equiv \lambda$

occurs only when $H_X(t) = \lambda t$, $S_X(t) = e^{-\lambda t}$
for Exponential random variable X

(ii) *Increasing hazard rate* = Aging, wearing-out

(iii) *Decreasing hazard rate* = ‘Burning-in’, mixture of exponential

Examples of Survival Hazards

- ‘Multi-hit model’ $X = V_1 + V_2 + \dots + V_r$ with indep. waiting times V_j for ‘shocks’, mutations, etc.

If V_j *iid* $\text{Expon}(\lambda)$, then $X \sim \text{Gamma}(r, \lambda)$
increasing-hazard if $r > 1$.

- ‘Mixture model’ $X \sim \text{Expon}(\tau)$, $\tau \sim G$ r.v.
Then can prove $h_X(t)$ decreasing : the idea is that individuals (X_i, τ_i) with higher τ_i die early !
- Weibull(λ, γ) power-law hazard $h(t) = \lambda \gamma t^{\gamma-1}$;
scale and power transformation of $V \sim \text{Expon}(1)$:
 $(V/\lambda)^{1/\gamma} \sim \text{Weib}(\lambda, \gamma)$ because:

$$S(t) = P((V/\lambda)^{1/\gamma} > t) = P(V > \lambda t^\gamma) = e^{-\lambda t^\gamma}$$

Hazard $h(t) \nearrow$ for $\gamma > 1$, \searrow for $\gamma < 1$

- *Bathtub-shaped* hazards in *Makeham* model:
 $h(t) = A + Be^{ct}$ ($A, B, c > 0$)

only if we add power-law term $\lambda \gamma t^{\gamma-1}$, $\gamma < 1$.

Pictures follow:

Parametric survival fcn with median 60

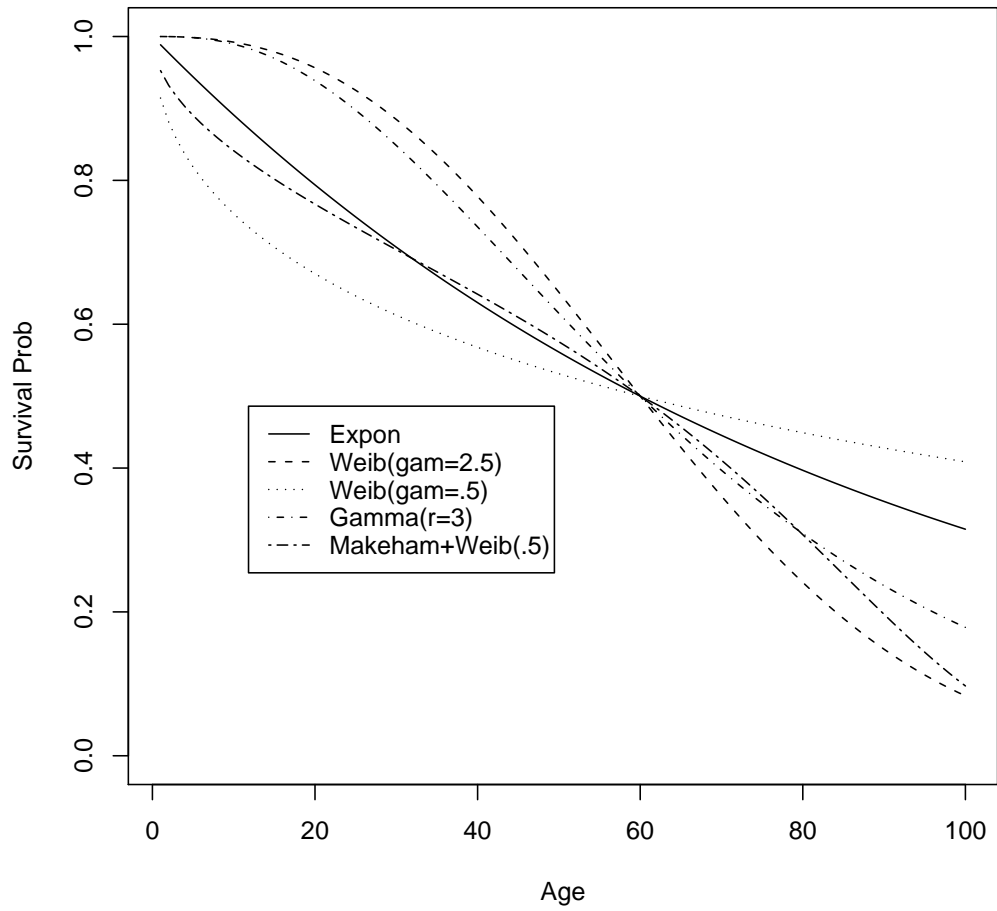


Figure 2: Graphs of survival functions from several parametric models designed to have common median 60.

Parametric Hazard fcn's with median at 60

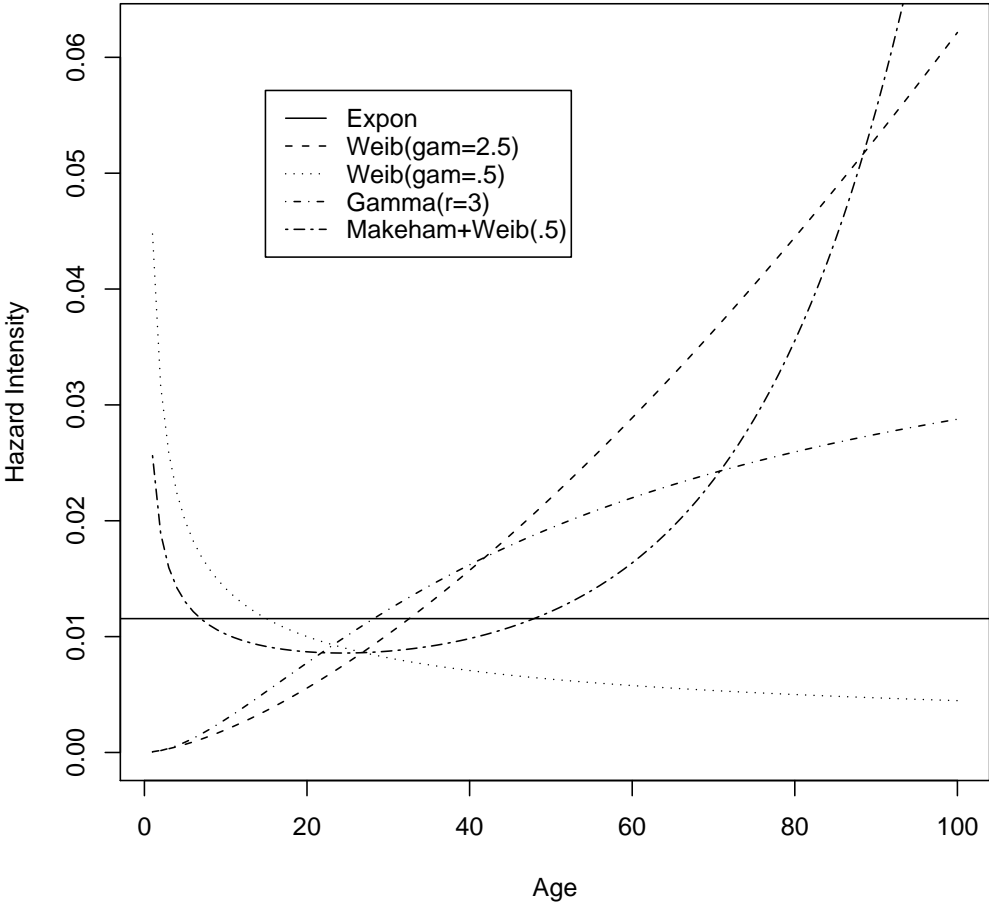


Figure 3: Graphs of cumulative hazard functions of several parametric models designed to have common median 60.