

## Chapter 2 Lecture in STAT 730, Spring 2026

Here: [Stochastic Regression](#) ideas for Shumway-Stoffer Chap. 2.

Previous lecture: stochastic regression via Partial Likelihood as in Slud & Kedem (1994, *Stat. Sinica*) and Kedem & Fokianos, **Regression Models in Time Series Analysis** (2002), especially conditionally specified Logistic Regression (+ other categorical time series models) with time-dependent covariates.

Topics in Ch. 2: classic linear regression models,  
search for trends and periodicities.

## Linear Regression as One-Step-Ahead Model

$$\text{Model is } \begin{pmatrix} Y_1 \\ Y_2 \\ \vdots \\ Y_n \end{pmatrix} = \mathbf{Z}\beta + \underline{\epsilon}, \quad \mathbf{Z} = \left( \begin{array}{c|c} \mathbf{X} & \mathbf{X}^* \\ \hline n \times p & n \times q \end{array} \right), \quad \beta = (\gamma, \gamma^*)$$

Slides **PLregrTalk2023.pdf** introduce **Partial Likelihood** idea.

OLS Regression: (treats  $\epsilon_t \stackrel{iid}{\sim} (0, \sigma^2)$  even if it isn't)

$$\log\text{PL} = \prod_{t=1}^n f_{Y_t | (Y_s, Z_s, s < t)}(\beta, \sigma^2) = -\frac{n}{2} \log(\sigma^2) - \frac{1}{2\sigma^2} \|\mathbf{Y} - \mathbf{Z}\beta\|^2$$

Assume that  $X_t, X_t^*$  processes are *adapted* to the timing of  $Y_t$  in the sense that  $(X_s, X_s^*, s \leq t)$  are (determined by  $\mathcal{F}_{t-1}$ ) observable up through time  $t - 1$

## OLS Estimation and Asymptotic Theory

Model:  $\epsilon_t = Y_t - \beta^{tr} \mathbf{Z}_t$ ,  $E(\epsilon_t | \mathcal{F}_{t-1}) = 0$ ,  $E(\epsilon_t^2 | \mathcal{F}_{t-1}) = \sigma^2$

$\sum_{s=1}^t \mathbf{Z}_s \epsilon_s$  is a martingale *recall def'n*

Conditional Estimating Equation  $\mathbf{Z}^{tr} (\mathbf{Y} - \mathbf{Z}\hat{\beta}) = \underline{0}$

$$SSE = \|\mathbf{Y} - \mathbf{Z}\hat{\beta}\|^2 = \mathbf{Y}^{tr} (\mathbf{I} - \mathbf{Z}(\mathbf{Z}^{tr}\mathbf{Z})^{-1}\mathbf{Z})\mathbf{Y}$$

$$\hat{\beta} - \beta = (\mathbf{Z}^{tr}\mathbf{Z})^{-1} \mathbf{Z}^{tr} \epsilon$$

In stochastic-regression, RHS not 'causal' (coef. of  $\epsilon_t$  in  $\mathcal{F}_{t-1}$ )

but  $\frac{1}{n} \mathbf{Z}^{tr} \mathbf{Z} = \frac{1}{n} \sum_{t=1}^n (Z_t)^{\otimes 2}$  may  $\xrightarrow{P} V$ , e.g., to  $E(Z_1^{\otimes 2})$

Under conditions:  $\frac{1}{\sqrt{n}}(\hat{\beta} - \beta) \xrightarrow{\mathcal{D}} \mathcal{N}_{p+q}(\underline{0}, \sigma^2 V^{-1})$

## ANOVA-style Model Comparisons

Consider null hypothesis  $H_0 : \gamma^* = \underline{0}$  that  $\mathbf{X}^*$  can be dropped

Analyze under  $H_0$ :  $SSE^{(r)} = \|\mathbf{Y} - \mathbf{X}\hat{\gamma}^{(r)}\|^2$

$^{(r)}$  stands for 'restricted' OLS under  $H_0$

and then consider incremental MSE =  $(\|\mathbf{Y} - \mathbf{Z}\hat{\beta}\|^2 - SSE^{(r)}) / (q\hat{\sigma}^{(r)2})$

leads to ideas of penalized logPL optimization via AIC, BIC

## Models with Deterministic Trend

**Example 1:** polynomial trends  $\mathbf{X} = \left( \mathbf{1} \mid \{t\}_{t=1}^n \right)$ ,  $\mathbf{X}^* = \{t^2\}_{t=1}^n$

Could formulate as regression, test if  $\mathbf{X}^*$  is necessary, **or**

**difference** once or twice:  $(I - B)Y_t \equiv Y_t - Y_{t-1}$ , **or**

$$(I - B)^2 Y_t \equiv Y_t - 2Y_{t-1} + Y_{t-2} \quad \text{and}$$

analyze as stationary process **under different assumptions on  $\epsilon_t$**

## Models Tailored to (Sin-Cosine) Bases

**Example 2:** sinusoidal trends, regress on

$$\mathbf{X} = \left( \{\cos(2\pi t\lambda_0)\}_{t=1}^n \mid \{\sin(2\pi t\lambda_0)\}_{t=1}^n \right)$$

to get coefficients  $\hat{\beta}_1(\lambda_0)$ ,  $\hat{\beta}_2(\lambda_0)$ .

*Frequency  $\lambda_0$  in place of period  $\omega_0 = 1/\lambda_0$*

When  $\lambda_0 = j/n$ , array of resulting coefficients  $\hat{\beta}_1(j, n)$ ,  $\hat{\beta}_2(j, n)$  are closely related to the **periodogram**  $\hat{\beta}_1^2(j/n) + \hat{\beta}_2^2(j/n)$  plotted against  $j/n$  as in Fig. 2.10.